

No asymptotics on the midterm (next Thursday)

Asymptotics: What happens as sample size goes to infinity?

$$\hat{\beta}_{OLS} = \underbrace{(\sum \mathbf{X}'\mathbf{X})^{-1}}_{1^{st} \text{ 5 weeks}} \underbrace{\sum \mathbf{X}'\mathbf{y}}_{\text{final 5 weeks}} \leftarrow \text{This is a fun of } n$$

\Rightarrow use notation $\hat{\beta}_n$. What happens to $\hat{\beta}_n$ as $n \rightarrow \infty$?

Defn: Let $\{a_n\}_{n=1}^{\infty} \in \mathbb{R}$. We say that $a_n \rightarrow a$ if $\forall \epsilon > 0$
 $\exists N(\epsilon) \in \mathbb{N}$ s.t. $\forall n \geq N, |a_n - a| < \epsilon$.

Defn: Let $\{f_n\}_{n=1}^{\infty}, f_n: \mathbb{R} \rightarrow \mathbb{R}$ be s.t. $\lim_{n \rightarrow \infty} f_n(x) = f(x) \quad \forall x \in \mathbb{R}$.

Then we say that $f_n \rightarrow f$ pointwise.

In probability, we have 4 different topologies:

Almost sure convergence

$$P(\omega: \lim_{n \rightarrow \infty} X_n(\omega) = X(\omega)) = 1 \iff X_n \xrightarrow{a.s.} X$$

$$\iff X_n - X \xrightarrow{a.s.} 0$$

Convergence in probability

$$\lim_{n \rightarrow \infty} P(\omega: \|X_n(\omega) - X(\omega)\| \leq \epsilon) = 1 \iff X_n \xrightarrow{P} X$$

$$= \lim_{n \rightarrow \infty} P_r(|X_n(\omega) - X(\omega)| \leq \epsilon) = 1 \quad (\text{univariate case})$$

$a.s. \Rightarrow P$ but $P \not\Rightarrow a.s.$

Defn: If $\hat{\theta}_n \xrightarrow{a.s.} \theta$, then $\hat{\theta}_n$ is strongly consistent.

Defn: If $\hat{\theta}_n \xrightarrow{P} \theta$, then $\hat{\theta}_n$ is (weakly) consistent.

Thm 6.1.10: If $X_n \xrightarrow{P} X$ and g is continuous a.e.,
 then $g(X_n) \xrightarrow{P} g(X)$. If $X_n \xrightarrow{a.s.} X$, then $g(X_n) \xrightarrow{a.s.} g(X)$.

Schemes for data:	iid	indep.	ident. distr.
	inid	indep.	not ident. distr.
	niid	not indep.	ident. distr.
	niid	not indep.	not ident. distr.

Here, we are only interested in LLN's for iid and inid schemes.

Laws of Large Numbers: (Strong, Weak)

$$\text{Denote } \bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$

$$\bar{\mu}_n = \frac{1}{n} \sum_{i=1}^n \mu_i \quad \text{where } \mu_i \equiv E[X_i] \text{ is true mean.}$$

(Kolmogorov)
iid LLN: Let $\{X_i\}$ be iid. Then $\bar{X}_n \xrightarrow{a.s.} E[X_i] \equiv \mu$
 $\Leftrightarrow \bar{X}_n - \mu \xrightarrow{a.s.} 0 \Leftrightarrow \bar{X}_n - \bar{\mu}_n \xrightarrow{a.s.} 0$

Proposition: Let g be a measurable function. If X_i and X_j are identically distributed, then $g(X_i)$ and $g(X_j)$ are identically distributed. If X_i and X_j are independent, then $g(X_i)$ and $g(X_j)$ are independent.

(Markov)

inid LLN: Let $\{X_i\}$ be independent with $\mu_i \equiv E[X_i] < +\infty \forall i$.

If for $\delta > 0$, $\sum_{i=1}^{\infty} \frac{E[|X_i - \mu_i|^{1+\delta}]}{\varepsilon^{1+\delta}} < +\infty$, then $\bar{X}_n - \bar{\mu}_n \xrightarrow{a.s.} 0$

Consistency of OLS

We are no longer in the classical model. New assumptions:

2.1. $Y_i = X_i \beta + \varepsilon_i$

2.2. $\{(X_i, \varepsilon_i)\}_{i=1}^n$ iid

2.3 (a) $E[X_i' \varepsilon_i] = 0$

(b) $E[|X_i \varepsilon_i|] < +\infty$

$\Leftrightarrow \{(X_i, \varepsilon_i)\}_{i=1}^n$ iid

X_i is uncorrelated with ε_i

(*) contemporaneous exogeneity

$\forall i \forall k$

- 2.4 (a) $E[|\sum_{ik} \epsilon_i|^2] < +\infty \quad \forall i=1, \dots, n, \quad \forall k$
 (b) $E[\sum_{i=1}^n \mathbf{X}_i' \mathbf{X}_i]$ is positive definite

Here, we are assuming random sampling.

$$\begin{aligned} \hat{\beta}_n &= \left(\frac{1}{n} \sum_{i=1}^n \mathbf{X}_i' \mathbf{X}_i \right)^{-1} \frac{1}{n} \sum_{i=1}^n \mathbf{X}_i' \mathbf{Y}_i \\ &= \left(\frac{1}{n} \sum_{i=1}^n \mathbf{X}_i' \mathbf{X}_i \right)^{-1} \frac{1}{n} \sum_{i=1}^n \mathbf{X}_i' (\mathbf{X}_i \beta + \epsilon_i) \\ &= \beta + \underbrace{\left(\frac{1}{n} \sum_{i=1}^n \mathbf{X}_i' \mathbf{X}_i \right)^{-1}}_{\text{by Kolmogorov}} \underbrace{\frac{1}{n} \sum_{i=1}^n \mathbf{X}_i' \epsilon_i}_{\text{by continuity thm}} \end{aligned}$$

by Kolmogorov $\xrightarrow{as.} (E[\sum_{i=1}^n \mathbf{X}_i' \mathbf{X}_i])^{-1}$ $\xrightarrow{as.} E[\sum_{i=1}^n \mathbf{X}_i' \epsilon_i] = 0$ by assumption

$$\Rightarrow \hat{\beta}_n \xrightarrow{as.} \beta + 0 = \beta$$

Therefore $\hat{\beta}_{OLS}$ is a consistent estimator