

Multivariate Delta Method  
(Greene, p.914, appendix D.)

If some vector of estimates,  $\hat{\theta}_1, \dots, \hat{\theta}_k$ , has the following asymptotic distribution

$$\sqrt{n} \left( \begin{bmatrix} \hat{\theta}_1 \\ \vdots \\ \hat{\theta}_k \end{bmatrix} - \begin{bmatrix} \theta_1 \\ \vdots \\ \theta_k \end{bmatrix} \right) \xrightarrow{d} N \left( \begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \Sigma \right)$$

where  $\theta_1, \dots, \theta_k$  is the true parameter and  $\Sigma$  is the  $k \times k$  asymptotic covariance matrix,

then some  $J$  continuous functions of the parameters,  $g_1(\theta_1, \dots, \theta_k), \dots, g_J(\theta_1, \dots, \theta_k)$ , have the following asymptotic distribution

$$\sqrt{n} \left( \begin{bmatrix} g_1(\hat{\theta}_1, \dots, \hat{\theta}_k) \\ \vdots \\ g_J(\hat{\theta}_1, \dots, \hat{\theta}_k) \end{bmatrix} - \begin{bmatrix} g_1(\theta_1, \dots, \theta_k) \\ \vdots \\ g_J(\theta_1, \dots, \theta_k) \end{bmatrix} \right) \xrightarrow{d} N \left( \begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix}, \Omega \right),$$

where  $\Omega$  is  $J \times J$  asymptotic covariance matrix.

$\Omega = \Psi \Sigma \Psi'$  and  $\Psi$  is the  $J \times k$  matrix of derivatives.

$$\Psi = \begin{bmatrix} \frac{\partial g_1(\hat{\theta}_1, \dots, \hat{\theta}_k)}{\partial \theta_1} & \dots & \frac{\partial g_1(\hat{\theta}_1, \dots, \hat{\theta}_k)}{\partial \theta_k} \\ \vdots & & \vdots \\ \frac{\partial g_J(\hat{\theta}_1, \dots, \hat{\theta}_k)}{\partial \theta_1} & \dots & \frac{\partial g_J(\hat{\theta}_1, \dots, \hat{\theta}_k)}{\partial \theta_k} \end{bmatrix}$$

An example using the normal distribution:  $z \sim N(\mu, \sigma^2)$

In section, we derived the MLE estimator of  $\mu$  and  $\sigma^2$  for a random sample of  $n$  observations.

$$\hat{\mu}_{ml} = 1/n \sum x_i \text{ and } \hat{\sigma}_{ml}^2 = 1/n \sum (x_i - \hat{\mu}_{ml})^2$$

We also derived the joint asymptotic distribution:

$$\sqrt{n} \left( \begin{bmatrix} \hat{\mu}_{ml} \\ \hat{\sigma}_{ml}^2 \end{bmatrix} - \begin{bmatrix} \mu \\ \sigma^2 \end{bmatrix} \right) \xrightarrow{d} N \left( \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma^2 & 0 \\ 0 & 2(\sigma^2)^2 \end{bmatrix} \right)$$

Let's assume we're actually interested in some other parameters,  $\eta_1$  and  $\eta_2$ . We know that these parameters are a function of the normal distribution parameters.

$$\eta_1 = g_1(\mu, \sigma^2) = 3\mu$$

$$\eta_2 = g_2(\mu, \sigma^2) = 5\left(\frac{\mu}{\sigma^2}\right)^2$$

By the Invariance Property of MLE, the MLE for  $\eta_1$  and  $\eta_2$  are

$$\hat{\eta}_{1,ml} = g_1(\hat{\mu}_{ml}, \hat{\sigma}_{ml}^2) = 3\hat{\mu}_{ml} = 3/n \sum x_i$$

$$\hat{\eta}_{2,ml} = g_2(\hat{\mu}_{ml}, \hat{\sigma}_{ml}^2) = 5\left(\frac{\hat{\mu}_{ml}}{\hat{\sigma}_{ml}^2}\right)^2 = 5\left(\frac{\sum x_i}{\sum (x_i - \hat{\mu}_{ml})^2}\right)^2$$

By the Delta Method, the joint asymptotic distribution of these two estimators is

$$\sqrt{n} \left( \begin{bmatrix} \hat{\eta}_{1,ml} \\ \hat{\eta}_{2,ml} \end{bmatrix} - \begin{bmatrix} \eta_1 \\ \eta_2 \end{bmatrix} \right) \xrightarrow{d} N \left( \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \Omega \right)$$

where in this case  $\Omega$  is 2x2 (JxJ)

$$\Omega = \begin{bmatrix} \frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} & \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} \\ \frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} & \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} \end{bmatrix} \begin{bmatrix} \sigma^2 & 0 \\ 0 & 2(\sigma^2)^2 \end{bmatrix} \begin{bmatrix} \frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} & \frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} \\ \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} & \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} \end{bmatrix}$$

$$= \begin{bmatrix} \frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} \sigma^2 & \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} 2(\sigma^2)^2 \\ \frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} \sigma^2 & \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} 2(\sigma^2)^2 \end{bmatrix} \begin{bmatrix} \frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} & \frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} \\ \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} & \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} \end{bmatrix}$$

$$= \begin{bmatrix} \frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} \sigma^2 \frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} + \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} 2(\sigma^2)^2 \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} & \text{con't} \\ \frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} \sigma^2 \frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} + \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} 2(\sigma^2)^2 \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} & \text{con't} \end{bmatrix}$$

$$\begin{bmatrix} \text{con't} & \frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} \sigma^2 \frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} + \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} 2(\sigma^2)^2 \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} \\ \text{con't} & \frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} \sigma^2 \frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} + \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} 2(\sigma^2)^2 \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} \end{bmatrix}$$

Computing the derivatives:

$$\frac{\partial g_1(\mu, \sigma^2)}{\partial \mu} = 3 \text{ and } \frac{\partial g_1(\mu, \sigma^2)}{\partial \sigma^2} = 0$$

$$\frac{\partial g_2(\mu, \sigma^2)}{\partial \mu} = \frac{10\mu}{\sigma^4} \text{ and } \frac{\partial g_2(\mu, \sigma^2)}{\partial \sigma^2} = -10 \frac{\mu^2}{\sigma^6}$$

$$\Omega = \begin{bmatrix} 9\sigma^2 & 3\sigma^2 \frac{10\mu}{\sigma^4} \\ 3\sigma^2 \frac{10\mu}{\sigma^4} & \frac{100\mu^2}{\sigma^8} \sigma^2 + 100 \frac{\mu^4}{\sigma^{12}} 2(\sigma^2)^2 \end{bmatrix}$$

Note: Although there is no correlation between the MLE estimates of  $\mu$  and  $\sigma^2$  ( $\text{cov}(\hat{\mu}_{ml}, \hat{\sigma}_{ml}^2) = 0$ ), there is correlation between the MLE estimates of  $\eta_1$  and  $\eta_2$ .

This is because the MLE estimates of these parameters are functions of the same random variable, in this case  $\hat{\mu}_{ml}$ .