

Quantile Regression (Theory):

Let \underline{Y} be a random variable. The τ -quantile of \underline{Y} is a number $Q_{\underline{Y}}(\tau)$ s.t.

$$\Pr[\underline{Y} \leq Q_{\underline{Y}}(\tau)] = \tau \quad (\text{continuous } \underline{Y} \text{ case}).$$

More generally, τ -quantile is defined as

$$Q_{\underline{Y}}(\tau) = \inf \{y: \Pr[\underline{Y} \leq y] \geq \tau\}$$

(Clearly, $Q_{\underline{Y}}(\tau)$ is the inverse of $F_{\underline{Y}}(y) = \Pr[\underline{Y} \leq y]$)

Skorohod representation (Monte-Carlo) representation of random variable \underline{Y} :

$$\underline{Y} = Q_{\underline{Y}}(U) \quad \text{where } U \sim \underline{U}[0, 1]$$

$$\Pr[\underline{Y} \leq y] = \Pr[Q_{\underline{Y}}(U) \leq y]$$

$$= \Pr[U \leq Q_{\underline{Y}}^{-1}(y)]$$

$$= \Pr[U \leq F_{\underline{Y}}(y)] = F_{\underline{Y}}(y)$$

Conditional (Regression) case:

$Q_{\underline{Y}}(\tau | \underline{X})$ solves $\Pr[\underline{Y} \leq Q_{\underline{Y}}(\tau | \underline{X}) | \underline{X}] = \tau$ so

that $Q_{\underline{Y}}(\tau | \underline{X})$ can be seen as the inverse of the conditional distribution function

$$F_{\underline{Y}}(y | \underline{X}) = \Pr[\underline{Y} \leq y | \underline{X}]$$

• Skorohod representation in the conditional case:

$$Y = Q_Y(U|X) \quad \text{where } U|X \sim U[0,1]$$

• We will use approximation

$$Q_Y(\tau|X) \approx X'\beta(\tau)$$

The corresponding regression equation is

$$Y = X'\beta(U), \quad U|X \sim U[0,1]$$

• random coefficient model. (very nonlinear)

• people differ in their propensity to consume

How does quantile regression compare to conventional regression?

QR versus Conventional Regression

Location model:

$$Y_i = X_i'\beta_0 + \varepsilon_i, \quad \varepsilon_i \perp X_i \quad \text{traditional model}$$

$$= X_i'\beta_0 + Q_{\varepsilon_i}(U_i)$$

$$U_i \sim U[0,1]$$

• slopes are constant, intercepts are random

• This is a special case of the above random coefficient model.

Has conditional quantile function

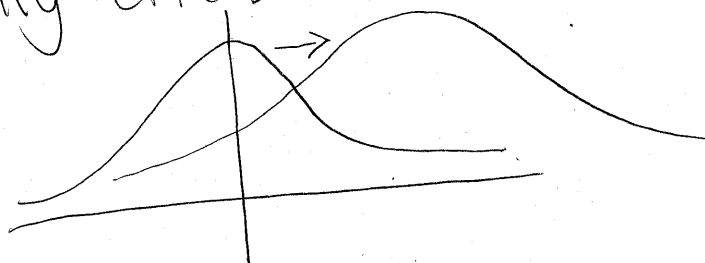
$$Q_Y(\tau|X) = X'\beta_0 + Q_{\varepsilon}(\tau)$$

$$= \mathbf{X}' \boldsymbol{\beta}(\tau), \text{ where } \boldsymbol{\beta}(\tau) = \begin{bmatrix} \beta_{01} + Q_{\varepsilon}(\tau) \\ \beta_{02} \\ \vdots \\ \beta_{0k} \end{bmatrix}$$

Location-Scale Model

$$Y_i = \mathbf{X}_i' \boldsymbol{\beta}_0 + (\mathbf{X}_i' \boldsymbol{\gamma}_0) \cdot \varepsilon_i \quad \varepsilon_i \perp \mathbf{X}_i$$

◦ only affect scale - not skewness/kurtosis



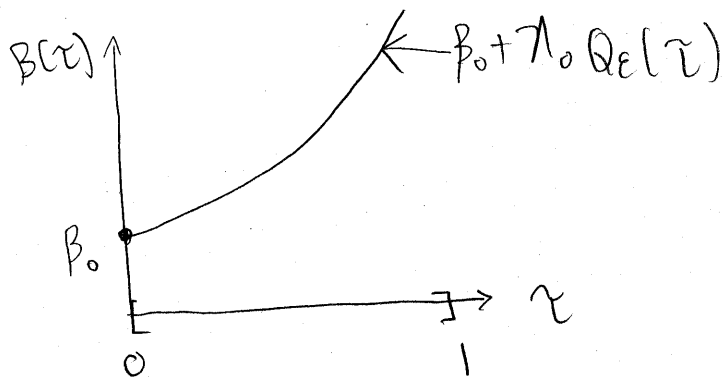
$$\Rightarrow Y_i = \mathbf{X}_i' \boldsymbol{\beta}_0 + (\mathbf{X}_i' \boldsymbol{\gamma}_0) \cdot Q_{\varepsilon}(U_i)$$

$$\Rightarrow Q_Y(\tau | \mathbf{X}) = \mathbf{X}' \boldsymbol{\beta}_0 + (\mathbf{X}' \boldsymbol{\gamma}_0) Q_{\varepsilon}(\tau)$$

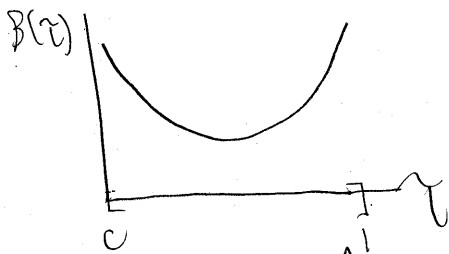
$$= \mathbf{X}' \boldsymbol{\beta}(\tau)$$

$$= \mathbf{X}' (\boldsymbol{\beta}_0 + \boldsymbol{\gamma}_0 Q_{\varepsilon}(\tau))$$

$$\text{where } \boldsymbol{\beta}(\tau) = \begin{bmatrix} \beta_{01} + \gamma_{01} Q_{\varepsilon}(\tau) \\ \beta_{02} + \gamma_{02} Q_{\varepsilon}(\tau) \\ \vdots \\ \beta_{0k} + \gamma_{0k} Q_{\varepsilon}(\tau) \end{bmatrix}$$



◦ This is restrictive, since it imposes linear structure and hence monotonicity. We reject the location-scale model if we have:



• This is inconsistent with the location-scale model.

$$\left[\begin{array}{l} \hat{\epsilon}_i = \frac{Y_i - X_i' \hat{\beta}}{X_i' \hat{\gamma}} \\ \hat{Q}_Y(z | X) = \beta_0(z) + (X' \hat{\gamma}) \hat{Q}_\epsilon(z) \end{array} \right] \leftarrow \text{side}$$

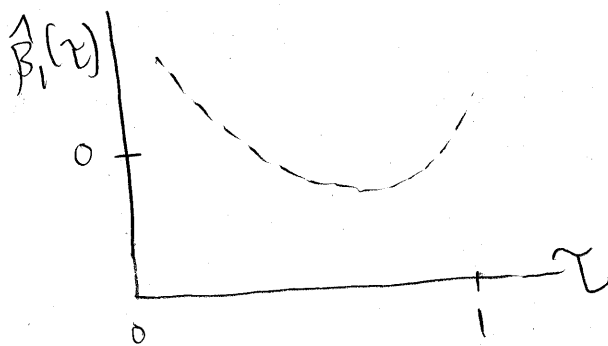
Doksum (1974)

- Annals of Statistics
- Biometrika

Y_i = duration of survival of guinea pigs

$X_i = \begin{bmatrix} 1 \\ D_i \end{bmatrix}$, D_i = indicator of vaccine treatments

$$Q_Y(z | X) = \beta_0(z) + \beta_1(z) \cdot D$$



z - "proneness"
◦ unobserved health or immunity.

$$Q_Y(z | \text{control}) = \beta_0(z)$$

$$Q_Y(z | \text{treatment}) = \beta_0(z) + \beta_1(z)$$

$$\Delta = \beta_1(z)$$