

$$R(\theta, w(\cdot)) = E_{\theta} [L(\theta - w(X))] \\ = \int [L(\theta - w(x))] f_X(x|\theta) dx$$

## Large sample criteria

### 2] Consistency

Defn 10.1.1: A sequence of estimators  $\{W_n\}$  is consistent for  $\theta$  if  $\text{plim } W_n = \theta \quad \forall \theta \in \Theta$ .

### Remarks:

1] As with  $\xrightarrow{P}$  in general, this just captures the idea of "close."

2] Key difference:  $\forall \theta \in \Theta$

3] Sense of closeness depends on method of convergence.

4] Recall from Chebyshev, if  $\overbrace{E_{\theta}[(W_n - \theta)^2]}^{\text{MSE}} \rightarrow 0$   
 $\forall \theta \in \Theta$ , then  $\text{plim } W_n = \theta$

$$\bullet E_{\theta}[(W_n - \theta)^2] = \text{Var}_{\theta}(W_n) + [E[W_n] - \theta]^2$$

$\bullet$  consistency is implied by  $\text{Var}_{\theta}(W_n) \rightarrow 0$   
 and  $\text{Bias}_{\theta} \rightarrow 0$ .

5] Thm 10.1.5: If  $\{W_n\}$  is consistent for  $\theta$  and an  $u$   
 a sequence s.t.  $a_n \rightarrow 1$  and  $b_n \rightarrow 0$ , then

$U_n = a_n W_n + b_n$  is consistent for  $\theta$ .

$\bullet$  Given one consistent estimator, there are  
 a continuum of consistent estimators.

6]  $W_n$  consistent for  $\theta$ ,

$$U_n = \begin{cases} B_n & n \leq 10,000 \\ W_n & n > 10,000 \end{cases}$$

$U_n$  is also consistent for  $\theta$

7] Thm 10.1.6: MLEs under some regularity conditions are consistent.

$$\hat{\theta}_{MLE}(x) = \operatorname{argmax}_{\theta \in \Theta} L_n(\theta|x)$$

• Need  $\Theta$  compact,  $L_n$  continuous in  $\theta$ ,  $L_n$  measurable in  $\mathcal{X}$ , and  $L_n \rightarrow L$  uniformly in  $\theta$ .

8] Method of moments:  $\underline{m} = g(\hat{\theta}) \Rightarrow \hat{\theta} = g^{-1}(m)$   
vector of moments

• if  $g^{-1}$  is nice enough, MM estimators are consistent.

eg:  $X_i \sim \text{Bernoulli}(p)$ . Let  $W_n(\mathcal{X}) = \bar{X}_n$ . Then

$$E[\bar{X}_n - p] = 0 \quad \forall p$$

$$\operatorname{Var}(\bar{X}_n - p) = \frac{p(1-p)}{n} \quad \forall p$$

Since  $\operatorname{Var}(\bar{X}_n) \rightarrow 0$ , and bias = 0  $\forall n$ , we have that  $\text{plim } \bar{X}_n = p$

9] Asymptotic Efficiency

Defn 10.1.1 A sequence of estimators  $\{W_n\}$  is asymptotically efficient for  $\tau(\theta)$  if

$$\sqrt{n} (W_n - \tau(\theta)) \xrightarrow{d} N(0, v(\theta)), \text{ where}$$

$$v(\theta) = \frac{(\tau'(\theta))^2}{E_{\theta} \left[ \left( \frac{\partial}{\partial \theta} \log f(x|\theta) \right)^2 \right]}$$

i.e.  $\tau(\theta) = \theta \Rightarrow (\tau'(\theta))^2 = 1^2 = 1$

$W_n$  has CLT and  $\text{Var} = \frac{1}{I(\theta)}$  (CRLB attained

asymptotically.)

e.g.  $\bar{X}_1, \dots, \bar{X}_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$

Let  $W_n = \bar{X}_n$ ,  $\text{Var}(\bar{X}_n) = \frac{p(1-p)}{n}$

$$\frac{\partial^2}{\partial p^2} \log L(p|x) = -\frac{n\bar{X}_n}{p^2} - \frac{n(1-\bar{X}_n)}{(1-p)^2}$$

$$\Rightarrow E \left[ -\frac{\partial^2}{\partial p^2} \log L(p|x) \right] = \frac{n}{p} + \frac{n}{1-p} = \frac{n}{p(1-p)} = I_n(p)$$

$$\text{CLT: } \sqrt{n} (\bar{X}_n - p) \xrightarrow{d} N(0, p(1-p))$$

Remarks:

1] Locally asymptotically normal.

2] iid Normal - typically get CRLB for all  $n$ .

Hypothesis Testing

Have a density  $f(x|\theta)$ . This is your economic model. ( $\theta \in \Theta$ )

A theory should restrict  $\Theta$ , i.e.  $\Theta \in \Theta_0 \not\subseteq \Theta$ .

Defn: A statistical hypothesis is a statement which implies that the true model belongs to a subset of possible models:

$$\Theta = \Theta_0 \cup \Theta_1 \quad \text{and} \quad \Theta_0 \cap \Theta_1 = \emptyset.$$

$\Theta \in \Theta_0 =$  null hypothesis

$H_0$

$\Theta \in \Theta_1 =$  alternative hypothesis

$H_1$  or  $H_A$

If  $\Theta_0 = \{\theta_0\}$  (point), we refer to this as a point null, or simple null.

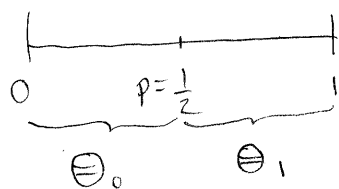
If  $\Theta_1 = \{\theta_1\}$ , we refer to this as a simple alternative.

Defn: A hypothesis testing procedure is a rule which specifies:

1] For which sample values we choose  $H_0$

2] For which sample values we choose  $H_1$ .

e.g.  $\bar{X}_1, \dots, \bar{X}_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$



$$H_0: p \leq \frac{1}{2}$$

$$H_1: p > \frac{1}{2}$$

Say we use  $\bar{X}_n$ . "If  $\bar{X}_n \leq \frac{1}{2}$ , choose  $H_0$ .

If  $\bar{X}_n > \frac{1}{2}$ , choose  $H_1$ ." This is a hypothesis testing procedure.

The sampling distribution of  $\bar{X}_n$  leads to classification errors.

Rejection region:  $\{(x_1, \dots, x_n) : \text{we choose } H_1\} \equiv R$

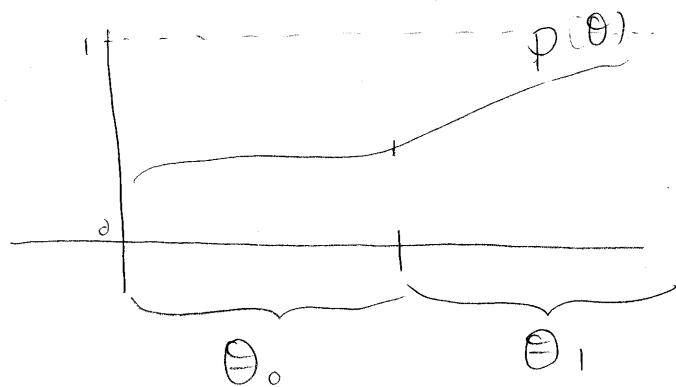
Acceptance region:  $\{(x_1, \dots, x_n) : \text{we choose } H_0\} \equiv A$

• We are going to make errors.

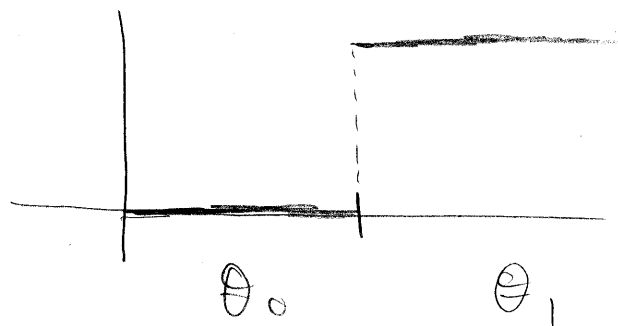
$$\Pr[\text{Type 1 error}] = \Pr[\bar{X} \in R \mid \theta \in \theta_0] = \alpha(\theta)$$

$$\Pr[\text{Type 2 error}] = \Pr[\bar{X} \in A \mid \theta \in \theta_1] = 1 - \beta$$

$$\text{Power} = \Pr[\bar{X} \in R \mid \theta] = p(\theta)$$



Ideal test:



Classical approach: want to increase  $p(\theta) \forall \theta \in \theta_1$  and decrease  $p(\theta) \forall \theta \in \theta_0$ .

1] Fix "tolerable" level of type I error

- size
- level

2] Maximize power given size.